

**G12 Mathematics Extended Part**  
**Module 1 – Calculus and Statistics**  
**Mock Examination**  
**Marking Scheme**

**Date: January 2023**

**Section A (50 marks)**

1. (a) Since  $E(X^2) = 4.3$ ,

$$\begin{aligned}(-2)^2(0.1) + (0)^2(0.2) + (1)^2(0.3) + a^2(0.4) &= 4.3 && 1M \\0.4a^2 + 0.7 &= 4.3 \\a^2 &= 9 \\a &= \pm 3\end{aligned}$$

Since  $a > 0$ ,  $a = 3$ . 1A

1. (b) Method 1

$$\begin{aligned}E(X) &= (-2)(0.1) + (0)(0.2) + (1)(0.3) + (3)(0.4) \\&= 1.3 \\E(3 - 2X) &= 3 - 2E(X) && * \\&= 3 - 2(1.3) \\&= 0.4 && 1A \\Var(X) &= E(X^2) - (E(X))^2 \\&= 4.3 - 1.3^2 && 1M \\&= 2.61 \\Var(3 - 2X) &= (-2)^2 Var(X) && * \\&= 4(2.61) \\&= 10.44 && 1A\end{aligned}$$

\*1M for either

Method 2

$$\begin{aligned} & E(3-2X) \\ &= [3-2(-2)](0.1) + [3-2(0)](0.2) + [3-2(1)](0.3) + [3-2(3)](0.4) \end{aligned}$$

\*\*

1A

= 0.4

Method 2.1

$$\begin{aligned} & E((3-2X)^2) \\ &= [3-2(-2)]^2(0.1) + [3-2(0)]^2(0.2) + [3-2(1)]^2(0.3) + [3-2(3)]^2(0.4) \end{aligned}$$

\*\*

= 10.6

$$\begin{aligned} & \text{Var}(3-2X) \\ &= E((3-2X)^2) - (E(3-2X))^2 \\ &= 10.6 - 0.4^2 \end{aligned}$$

1M

1A

\*\*1M for either

= 10.44

Method 2.2

$$\begin{aligned} & \text{Var}(3-2X) \\ &= E([(3-2X) - E(3-2X)]^2) \\ &= E([(3-2X) - 0.4]^2) \end{aligned}$$

1M

$$\begin{aligned} &= E((2.6-2X)^2) \\ &= [2.6-2(-2)]^2(0.1) + [2.6-2(0)]^2(0.2) + [2.6-2(1)]^2(0.3) + [2.6-2(3)]^2(0.4) \end{aligned}$$

\*\*

= 10.44

1A

\*\*1M for either

$$\begin{aligned}
2. \quad (a) \quad & P(A \cap B) \\
& = P(A|B) \cdot P(B) \\
& = \frac{1}{6} P(B)
\end{aligned}$$

1A

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

$$\frac{9}{20} = \frac{1}{5} + P(B) - \frac{1}{6} P(B)$$

1M

$$\frac{1}{4} = \frac{5}{6} P(B)$$

$$P(B) = \frac{3}{10}$$

1A

2. (b) Method 1

$$\begin{aligned}
& P(A' \cap B) \\
& = P(B) - P(A \cap B)
\end{aligned}$$

$$= \frac{3}{10} - \frac{1}{6} \left( \frac{3}{10} \right)$$

1M

$$= \frac{1}{4}$$

1A

Method 2

$$\begin{aligned}
& P(A' \cap B) \\
& = P(A \cup B) - P(A)
\end{aligned}$$

$$= \frac{9}{20} - \frac{1}{5}$$

1M

$$= \frac{1}{4}$$

1A

2. (c)

Method 1

Note that  $A' \cap B' = (A \cup B)'$ .

$$\begin{aligned} P(A' \cap B') &= 1 - P(A \cup B) \\ &= 1 - \frac{9}{20} \\ &= \frac{11}{20} \end{aligned}$$

Method 2

$$\begin{aligned} P(A' \cap B') &= P(A') - P(A' \cap B) \\ &= \left(1 - \frac{1}{5}\right) - \frac{1}{4} \\ &= \frac{11}{20} \end{aligned}$$

Since  $P(A' \cap B') \neq 0$ ,  $A'$  and  $B'$  are not mutually exclusive.

1A f.t.

3. (a) Since  $P(X < 10) = 0.8413$ ,

$$P\left(Z < \frac{10 - \mu}{2}\right) = 0.8413 \quad 1M$$

$$P\left(0 < Z < \frac{10 - \mu}{2}\right) = 0.8413 - 0.5$$

$$P\left(0 < Z < \frac{10 - \mu}{2}\right) = 0.3413$$

$$\frac{10 - \mu}{2} = 1$$

$$\mu = 8 \quad 1A$$

3. (b) (i)  $P(A)$

$$= P(5 \leq X \leq 6)$$

$$= P\left(\frac{5-8}{2} \leq Z \leq \frac{6-8}{2}\right)$$

$$= P(-1.5 \leq Z \leq -1)$$

$$= 0.4332 - 0.3413$$

$$= 0.0919 \quad 1A$$

3. (b) (ii)  $P(B_t)$

$$= P(-1 < X - t < 1)$$

$$= P(t - 1 < X < t + 1)$$

Since  $X$  is a normal random variable,  $P(B_t)$  attains its greatest value at  $t = \mu = 8$  and  $T = 8$ . 1M

Then  $B_T$  is the event that

$$T - 1 < X < T + 1$$

$$8 - 1 < X < 8 + 1$$

$$7 < X < 9$$

It follows that  $A \cap B_T = \emptyset$  and  $P(A \cap B_T) = 0$ . 1M

Note that  $P(A) > 0$  and  $P(B_T) > 0$ .

Since  $P(A \cap B_T) \neq P(A)P(B_T)$ ,  $A$  and  $B_T$  are not independent.

1A ft.

4. (a) Method 1

$$\begin{aligned} &\text{The required probability} && 1M \\ &= (0.6)(1-0.7) + (0.3)(1-0.8) + (0.1)(1-0.9) && 1A \\ &= 0.25 \end{aligned}$$

Method 2

$$\begin{aligned} &\text{The required probability} && 1M \\ &= 1 - [(0.6)(0.7) + (0.3)(0.8) + (0.1)(0.9)] && 1A \\ &= 0.25 \end{aligned}$$

4. (b) (i) The required probability

$$\begin{aligned} &= C_3^5 (0.25)^3 (1-0.25)^2 + C_4^5 (0.25)^4 (1-0.25) + (0.25)^5 && 1M \\ &= \frac{53}{512} && 1A \\ &= 0.103515625 \\ &= 0.1035 \text{ (cor. to 4 d. p.)} \end{aligned}$$

4. (b) (ii) The probability that Harrods is late for school on at least 3 consecutive days in the week given that he is late for school on at least 3 days in the week

$$\begin{aligned} &= \frac{(3)(0.25)^3 (1-0.25)^2 + (4)(0.25)^4 (1-0.25) + (0.25)^5}{\frac{53}{512}} && 1M \text{ for numerator +} \\ & && 1M \text{ for denominator} \\ &= \frac{20}{53} \\ &\approx 0.37735849 \\ &< 0.4 \end{aligned}$$

The claim is disagreed. 1A ft.

5. (a)  $(4 - e^{-px})^3$

$$= 4^3 - C_1^3(4)^2(e^{-px}) + C_2^3(4)(e^{-px})^2 - (e^{-px})^3 \quad 1M$$

$$= 64 - 48e^{-px} + 12e^{-2px} - e^{-3px}$$

$$= 64 - 48 \left[ 1 + (-px) + \frac{(-px)^2}{2!} + \frac{(-px)^3}{3!} + \dots \right] \\ + 12 \left[ 1 + (-2px) + \frac{(-2px)^2}{2!} + \frac{(-2px)^3}{3!} + \dots \right] \\ - \left[ 1 + (-3px) + \frac{(-3px)^2}{2!} + \frac{(-3px)^3}{3!} + \dots \right] \quad \left. \vphantom{\begin{matrix} 64 \\ 48 \\ 12 \\ - \end{matrix}} \right\} 1M$$

$$= 27 + 27px - \frac{9p^2}{2}x^2 - \frac{7p^3}{2}x^3 + \dots \quad 1A$$

5. (b)

$$\left( -\frac{9p^2}{2} \right) + \left( -\frac{7p^3}{2} \right) = -8$$

$$7p^3 + 9p^2 - 16 = 0$$

$$(p-1)(7p^2 + 16p + 16) = 0$$

$$p = 1 \text{ or } 7p^2 + 16p + 16 = 0$$

Consider  $7p^2 + 16p + 16 = 0$ .

$$\Delta \\ = 16^2 - 4(7)(16)$$

$$= -192$$

$$< 0$$

It follows that  $7p^2 + 16p + 16 = 0$  has no real roots.

Thus  $p = 1$ .

1M

1A

6. (a) Since  $r$  is an integer,  $x^{2r} = (x^2)^r$  for all non-zero real values of  $x$ .

$$\text{Then } g(x) = \frac{\ln x^{2r}}{\ln x^2 + 1} = \frac{\ln(x^2)^r}{\ln x^2 + 1} = \frac{r \ln x^2}{\ln x^2 + 1} \text{ for all non-zero real values of } x.$$

$$\lim_{x \rightarrow \infty} g(x) = 3$$

$$\lim_{x \rightarrow \infty} \frac{r \ln x^2}{\ln x^2 + 1} = 3$$

$$\lim_{x \rightarrow \infty} \frac{\frac{r \ln x^2}{\ln x^2}}{\frac{\ln x^2 + 1}{\ln x^2}} = 3$$

$$\lim_{x \rightarrow \infty} \frac{r}{1 + \frac{1}{\ln x^2}} = 3 \quad 1M$$

$$\frac{r}{1+0} = 3$$

$$r = 3 \quad 1$$

6. (b) Note that  $g(x) = \frac{3 \ln x^2}{\ln x^2 + 1}$  for all non-zero real values of  $x$ .

$$g'(x) = \frac{3 \left[ \left( \frac{2x}{x^2} \right) (\ln x^2 + 1) - (\ln x^2) \left( \frac{2x}{x^2} \right) \right]}{(\ln x^2 + 1)^2} \quad 1M$$

$$= \frac{6}{x(\ln x^2 + 1)^2}$$

The equation of the tangent to the curve  $y = g(x)$  at  $x = e$  is

$$\frac{y - g(e)}{x - e} = g'(e)$$

$$\frac{y - \frac{3 \ln e^2}{\ln e^2 + 1}}{x - e} = \frac{6}{e(\ln e^2 + 1)^2} \quad 1M$$

$$\frac{y - 2}{x - e} = \frac{2}{3e}$$

$$3ey - 6e = 2x - 2e$$

$$2x - 3ey + 4e = 0 \quad 1A$$



7. (a)  $h'(x)$

$$= 2x + 14 - \frac{72}{x^2}$$

1M

Put  $h'(x) = 0$ .

$$2x + 14 - \frac{72}{x^2} = 0$$

$$x^3 + 7x^2 - 36 = 0$$

$$(x-2)(x^2 + 9x + 18) = 0$$

1M

$$(x-2)(x+3)(x+6) = 0$$

$$x = 2 \text{ (rej.) or } -3 \text{ or } -6$$

Method 1 [First Derivative Test]

$x$	$x = -7$	$-7 < x < -6$	$x = -6$	$-6 < x < -3$	$x = -3$
$h(x)$	$-\frac{415}{7}$		$-60$		$-57$
$h'(x)$	$-$	$-$	$0$	$+$	$0$

$x$	$-3 < x < -1$	$x = -1$
$h(x)$		$-85$
$h'(x)$	$-$	$-$

Method 2 [Second Derivative Test]

$$h''(x)$$

$$= 2 - 72 \left( -\frac{2}{x^3} \right)$$

$$= 2 + \frac{144}{x^3}$$

$$h''(-6)$$

$$= 2 + \frac{144}{(-6)^3}$$

$$= \frac{4}{3}$$

$$> 0$$

$$h''(-3)$$

$$= 2 + \frac{144}{(-3)^3}$$

$$= -\frac{10}{3}$$

$$< 0$$

Note that  $h(-7) = -\frac{415}{7}$ ,  $h(-6) = -60$ ,  $h(-3) = -57$  and  $h(-1) = -85$ .

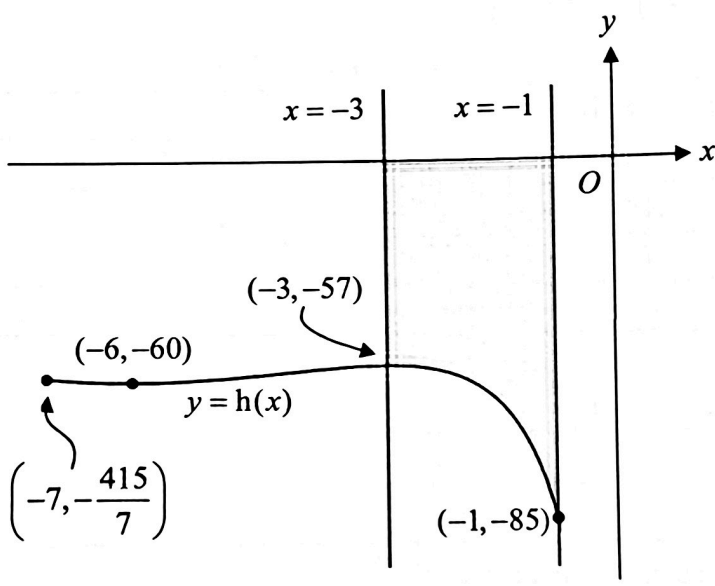
1M for first/second derivative test

The greatest value of  $h(x)$  is attained at  $x = -3$  and the least value of  $h(x)$  is attained at  $x = -1$ .

1A for both

Thus  $\alpha = -3$  and  $\beta = -1$ .

7. (b)



The required area

$$= -\int_{\alpha}^{\beta} h(x) dx$$

$$= -\int_{-3}^{-1} \left( x^2 + 14x + \frac{72}{x} \right) dx$$

1M for sign and limits

$$= -\left[ \frac{1}{3}x^3 + 14\left(\frac{1}{2}x^2\right) + 72\ln|x| \right]_{-3}^{-1}$$

1M

$$= -\left[ \frac{1}{3}x^3 + 7x^2 + 72\ln|x| \right]_{-3}^{-1}$$

$$= \frac{142}{3} + 72\ln 3$$

1A

8. (a)  $\frac{d}{dx}(x^2 e^{kx^2})$

$$= (2x)(e^{kx^2}) + (x^2)(2kxe^{kx^2}) \quad 1M$$

$$= 2xe^{kx^2} + 2kx^3 e^{kx^2}$$

Let  $u = kx^2$ . Then  $du = 2kx dx$ .

It follows that

$$2kx^3 e^{kx^2} = \frac{d}{dx}(x^2 e^{kx^2}) - 2xe^{kx^2}$$

$$\int 2kx^3 e^{kx^2} dx = x^2 e^{kx^2} - \int 2xe^{kx^2} dx \quad 1M$$

$$\int 2kx^3 e^{kx^2} dx = x^2 e^{kx^2} - \frac{1}{k} \int e^u du \quad 1M$$

$$2k \int x^3 e^{kx^2} dx = x^2 e^{kx^2} - \frac{1}{k} e^u + C'$$

$$\int x^3 e^{kx^2} dx = \frac{1}{2k} x^2 e^{kx^2} - \frac{1}{2k^2} e^{kx^2} + C \quad 1A$$

8. (b) (i) Since  $\frac{dy}{dx} = x^3 e^{kx^2}$ ,

$$y = \int x^3 e^{kx^2} dx$$

$$y = \frac{1}{2k} x^2 e^{kx^2} - \frac{1}{2k^2} e^{kx^2} + C$$

Put  $\left(0, 1 - \frac{1}{2k^2}\right)$  into the equation.

$$1 - \frac{1}{2k^2} = \frac{1}{2k} (0)^2 e^{k(0)^2} - \frac{1}{2k^2} e^{k(0)^2} + C \quad 1M$$

$$C = 1$$

It follows that

$$y = \frac{1}{2k} x^2 e^{kx^2} - \frac{1}{2k^2} e^{kx^2} + 1$$

$$y - 1 = \frac{e^{kx^2}}{2k^2} (kx^2 - 1)$$

$$\frac{y-1}{kx^2-1} = \frac{e^{kx^2}}{2k^2}$$

$$\ln\left(\frac{y-1}{kx^2-1}\right) = \ln\left(\frac{e^{kx^2}}{2k^2}\right)$$

$$\ln\left(\frac{y-1}{kx^2-1}\right) = kx^2 - \ln(2k^2) \quad 1A$$

8. (b) (ii) Note that the slope is  $k = \frac{1}{e^2}$ .

Method 1

When  $\ln\left(\frac{y-1}{kx^2-1}\right) = 0$ ,

$$kx^2 - \ln(2k^2) = 0$$

1M

$$\frac{x^2}{e^2} - \ln\left(\frac{2}{(e^2)^2}\right) = 0$$

$$\frac{x^2}{e^2} - \ln\left(\frac{2}{e^4}\right) = 0$$

$$x^2 = -(4 - \ln 2)e^2$$

This contradicts the fact that  $x^2 \geq 0$  for all real values of  $x$ .

Method 2

$$\ln\left(\frac{y-1}{kx^2-1}\right)$$

$$= kx^2 - \ln(2k^2)$$

$$= \frac{x^2}{e^2} - \ln\left(\frac{2}{(e^2)^2}\right)$$

$$= \frac{x^2}{e^2} - \ln\left(\frac{2}{e^4}\right)$$

$$\geq -\ln\left(\frac{2}{e^4}\right)$$

$$\left[ \because \frac{x^2}{e^2} \geq 0 \right]$$

1M

$$= 4 - \ln 2$$

$$> 0$$

It follows that the graph does not have any intercept on the horizontal axis.

The claim is incorrect.

1A ft.

**Section B (50 marks)**

9. (a) The required probability

$$= \frac{e^{-1}1^0}{0!} + \frac{e^{-1}1^1}{1!} + \frac{e^{-1}1^2}{2!}$$

$$= 2.5e^{-1}$$

$$\approx 0.919698602$$

$$= 0.9197 \text{ (cor. to 4 d. p.)}$$

1M for Poisson  
probability

1A

9. (b) (i) Note that the mean number of mistakes introduced by the secretary to each page of any document is  $\lambda$  and the mean number of mistakes introduced by the secretary to any document is  $5\lambda$ .

$$\frac{\left(\frac{e^{-1}1^2}{2!}\right)\left(\frac{e^{-5\lambda}(5\lambda)^0}{0!}\right)}{\left(\frac{e^{-1}1^0}{0!}\right)\left(\frac{e^{-5\lambda}(5\lambda)^2}{2!}\right) + \left(\frac{e^{-1}1^1}{1!}\right)\left(\frac{e^{-5\lambda}(5\lambda)^1}{1!}\right) + \left(\frac{e^{-1}1^2}{2!}\right)\left(\frac{e^{-5\lambda}(5\lambda)^0}{0!}\right)} = 0.16$$

1M for numerator + 1M for denominator

Award even if  $\lambda$  is used in place of  $5\lambda$

$$\frac{\frac{1}{2}}{\frac{25}{2}\lambda^2 + 5\lambda + \frac{1}{2}} = 0.16$$

$$\frac{1}{25\lambda^2 + 10\lambda + 1} = 0.16$$

$$25\lambda^2 + 10\lambda - 5.25 = 0$$

$$\lambda^2 + 0.4\lambda - 0.21 = 0$$

$$(\lambda - 0.3)(\lambda + 0.7) = 0$$

$$\lambda = 0.3 \text{ or } -0.7$$

Since  $\lambda \geq 0$ ,  $\lambda = 0.3$ .

1

9. (b) (ii) (1) The probability that a document prepared by Noah and processed by his secretary contains no mistakes

$$= \left( \frac{e^{-1} 1^0}{0!} \right) \left( \frac{e^{-1.5} 1.5^0}{0!} \right)$$

$$= e^{-2.5}$$

The probability that a document prepared by Noah and processed by his secretary contains exactly 1 mistake

$$= \left( \frac{e^{-1} 1^1}{1!} \right) \left( \frac{e^{-1.5} 1.5^0}{0!} \right) + \left( \frac{e^{-1} 1^0}{0!} \right) \left( \frac{e^{-1.5} 1.5^1}{1!} \right)$$

$$= 2.5e^{-2.5}$$

The probability that a document prepared by Noah and processed by his secretary contains exactly 2 mistakes

$$= \left( \frac{e^{-1} 1^2}{2!} \right) \left( \frac{e^{-1.5} 1.5^0}{0!} \right) + \left( \frac{e^{-1} 1^1}{1!} \right) \left( \frac{e^{-1.5} 1.5^1}{1!} \right) + \left( \frac{e^{-1} 1^0}{0!} \right) \left( \frac{e^{-1.5} 1.5^2}{2!} \right)$$

$$= 3.125e^{-2.5}$$

The probability that a document prepared by Noah and processed by his secretary is *accurate*

$$= e^{-2.5} + 2.5e^{-2.5} + 3.125e^{-2.5}$$

$$= 6.625e^{-2.5}$$

The required expected number of mistakes

$$= (0) \left( \frac{e^{-2.5}}{6.625e^{-2.5}} \right) + (1) \left( \frac{2.5e^{-2.5}}{6.625e^{-2.5}} \right) + (2) \left( \frac{3.125e^{-2.5}}{6.625e^{-2.5}} \right)$$

1M for conditional probability + 1M for expected value

$$= \frac{70}{53} \text{ or } 1.3208 \text{ (cor. to 4 d. p.)} \quad 1A$$

9. (b) (ii) (2) The required probability

$$= [C_2^5 (6.625e^{-2.5})^2 (1 - 6.625e^{-2.5})^3 (6.625e^{-2.5})] [C_2^4 (6.625e^{-2.5})^2 (1 - 6.625e^{-2.5})^2]$$

1M for first factor + 1M for second factor

$$\approx 0.056378512$$

$$= 0.0564 \text{ (cor. to 4 d. p.)}$$

1A

10. (a) (i) A 95% confidence interval for  $\mu$

$$= \left( 35.2 - 1.96 \left( \frac{2.5}{\sqrt{64}} \right), 35.2 + 1.96 \left( \frac{2.5}{\sqrt{64}} \right) \right) \quad 1M + 1A$$
$$= (34.5875, 35.8125) \quad 1A$$

10. (a) (ii) Let  $z$  be the critical value of the new confidence interval.

$$2z \left( \frac{2.5}{\sqrt{64}} \right) \leq 2(1.96) \left( \frac{2.5}{\sqrt{64}} \right) (1 - 20\%) \quad 1M$$
$$z \leq 1.568$$

The confidence level of the new confidence interval

$$\leq 2(0.4406)(100\%) \quad 1M$$
$$= 88.12\%$$

Thus  $0 < k \leq 88.12$ . 1A

10. (b) (i) Let  $b \text{ cm}^2$  be the sum of squares of the deviations of the lengths of the rods in sample  $B$  from the mean of the lengths of the rods in sample  $B$ .

$$\frac{b}{n} = \left( \frac{b}{n-1} \right) (1 - 20\%) \quad 1M$$
$$\frac{n-1}{n} = 0.98$$
$$n = 50 \quad 1$$

10. (b) (ii) (1)  $\bar{x}$  1M

$$= \frac{(35.2)(64) + 1840}{64 + 50}$$

$$\approx 35.90175439$$

\*

$$= 35.9018 \text{ (cor. to 4 d. p.)}$$

Let  $a \text{ cm}^2$  be the sum of squares of the lengths of the rods in sample  $A$ .

$$2.1^2 = \frac{a - (64)(35.2)^2}{64 - 1}$$

\*\*

$$a = 79576.39$$

The sum of squares of the lengths of the rods in sample  $A$  is  $79576.39 \text{ cm}^2$ .

$$s = \sqrt{\frac{(79576.39 + 68018.25) - (64 + 50)(35.90175439)^2}{(64 + 50) - 1}}$$

\*\*

$$\approx 2.409310203$$

$$= 2.4093 \text{ (cor. to 4 d. p.)}$$

\*

\*\*1M for either + \*1A for both

10. (b) (ii) (2) Let  $\bar{X} \text{ cm}$  be the mean of the lengths of the rods in sample  $D$ .

By the central limit theorem,  $\bar{X} \sim N\left(\bar{x}, \frac{s^2}{70}\right)$  approximately.

1M

The required probability

$$= P(70\bar{X} > 2500)$$

1A

$$\approx P\left(Z > \frac{\frac{2500}{70} - \bar{x}}{\frac{s}{\sqrt{70}}}\right)$$

$$\approx P\left(Z > \frac{\frac{2500}{70} - 35.90175439}{\frac{2.409310203}{\sqrt{70}}}\right)$$

$$\approx P(Z > -0.651006016)$$

$$\approx 0.5 + 0.2422$$

$$= 0.7422$$

1A



11. (a)  $f'(x)$

$$= \frac{(6x^5)(x^3-1) - (x^6)(3x^2)}{(x^3-1)^2} \quad 1M$$

$$= \frac{3x^5(x^3-2)}{(x^3-1)^2} \quad 1A$$

$$= \frac{3x^8 - 6x^5}{(x^3-1)^2}$$

$f''(x)$

$$= \frac{[3(8x^7) - 6(5x^4)](x^3-1)^2 - (3x^8 - 6x^5)[2(x^3-1)(3x^2)]}{(x^3-1)^4} \quad 1M$$

$$= \frac{6x^4(x^3-1)[(4x^3-5)(x^3-1) - (3x^6-6x^3)]}{(x^3-1)^4}$$

$$= \frac{6x^4(x^6-3x^3+5)}{(x^3-1)^3} \quad 1A$$

$$= \frac{6x^{10} - 18x^7 + 30x^4}{(x^3-1)^3}$$

11. (b) (i)  $I$

$$= \int_{-2}^0 f(x) dx$$

$$\approx \frac{1}{2} \left( \frac{0 - (-2)}{5} \right) \{f(-2) + 2[f(-1.6) + f(-1.2) + f(-0.8) + f(-0.4)] + f(0)\}$$

1M

$$\approx -3.247832838$$

$$= -3.2478 \text{ (cor. to 4 d. p.)}$$

1A

11. (b) (ii) Method 1

$$J$$

$$= \int_{-2}^0 \frac{4x^3 - 2}{x^3 - 1} dx$$

$$= -2 \int_{-2}^0 \frac{-2x^3 + 1}{x^3 - 1} dx$$

$$= -2 \int_{-2}^0 \frac{(x^6 - 2x^3 + 1) - x^6}{x^3 - 1} dx \quad 1M$$

$$= -2 \int_{-2}^0 \frac{(x^3 - 1)^2 - x^6}{x^3 - 1} dx$$

Method 2

Note that  $\frac{x^6}{x^3-1} = (x^3+1) + \frac{1}{x^3-1}$  and  $\frac{1}{x^3-1} = \frac{x^6}{x^3-1} - (x^3+1)$ .

$$\frac{4x^3-2}{x^3-1}$$

$$= 4 + \frac{2}{x^3-1}$$

$$= 4 + 2 \left[ \frac{x^6}{x^3-1} - (x^3+1) \right]$$

1M

$$= 2 - 2x^3 + \frac{2x^6}{x^3-1}$$

$J$

$$= \int_{-2}^0 \frac{4x^3-2}{x^3-1} dx$$

$$= \int_{-2}^0 \left( 2 - 2x^3 + \frac{2x^6}{x^3-1} \right) dx$$

$$= -2 \int_{-2}^0 (x^3-1) dx + 2 \int_{-2}^0 \frac{x^6}{x^3-1} dx$$

1M

$$= -2 \left[ \frac{1}{4} x^4 - x \right]_{-2}^0 + 2I$$

1M

$$\approx -2(-6) + 2(-3.247832838)$$

$$\approx 5.504334324$$

$$= 5.5043 \text{ (cor. to 4 d. p.)}$$

1A

11. (b) (iii) Note that  $f''(x) < 0$  for all  $-2 \leq x < 0$ .

The estimate in (b)(i) is an under-estimate.

1

From (b)(ii),

$$J = 12 + 2I$$

$$J > 12 + 2(-3.247832838)$$

$$J > 5.504334324$$

$$J > 5.5$$

The claim is correct.

1A f.t.

12. (a) Since  $T \geq 0$  and  $k > 1$ ,  $T+k > 1$  and  $\ln(T+k) > 0$ .

$$\begin{aligned} & \frac{d^2v}{dt^2} \\ &= \frac{300 \left\{ \left[ 2\ln(t+k) \right] \left( \frac{1}{t+k} \right) (t+k) - [\ln(t+k)]^2 (1) \right\}}{(t+k)^2} && 1M + 1A \\ &= \frac{300\ln(t+k)[2 - \ln(t+k)]}{(t+k)^2} \end{aligned}$$

Since the rate of change of the speed of the particle is maximum  $T$  s after the start of the experiment,

$$\begin{aligned} & \left. \frac{d^2v}{dt^2} \right|_{t=T} = 0 && 1M \\ & \frac{300\ln(T+k)[2 - \ln(T+k)]}{(T+k)^2} = 0 \\ & 2 - \ln(T+k) = 0 \\ & \ln(T+k) = 2 \\ & T+k = e^2 \\ & T = e^2 - k && 1A \end{aligned}$$

Remark: As the existence of a maximum value of  $\frac{dv}{dt}$  is stated in the

question and  $\left. \frac{d^2v}{dt^2} \right|_{t=T} = 0$  gives rise to only one possible value of

$T$ , testing for maximum is not necessary.

12. (b) (i) Method 1

$$\text{Let } u = \ln(t+k). \text{ Then } du = \frac{1}{t+k} dt. \quad 1M$$

$$\text{When } t=0, u = \ln k \text{ and } v = 100(\ln 2)^3.$$

$$\text{When } t=T, u = \ln(T+k) \text{ and } v = 800.$$

$$800 - 100(\ln 2)^3 = \int_0^T \frac{300[\ln(t+k)]^2}{t+k} dt \quad 1M$$

$$800 - 100(\ln 2)^3 = 300 \int_{\ln k}^{\ln(T+k)} u^2 du$$

$$8 - (\ln 2)^3 = 3 \left[ \frac{1}{3} u^3 \right]_{\ln k}^{\ln(T+k)} \quad 1M$$

$$8 - (\ln 2)^3 = 8 - (\ln k)^3$$

$$k = 2 \quad 1$$

Method 2

1M

Let  $u = \ln(t+k)$ . Then  $du = \frac{1}{t+k} dt$ .

$$\begin{aligned} v &= \int \frac{300[\ln(t+k)]^2}{t+k} dt \\ &= 300 \int u^2 du \\ &= 300 \left( \frac{1}{3} u^3 \right) + C \\ &= 100[\ln(t+k)]^3 + C \end{aligned}$$

1M

When  $t = T$ ,  $v = 800$ .

$$100[\ln(T+k)]^3 + C = 800$$

\*

$$100(2)^3 + C = 800$$

$$C = 0$$

When  $t = 0$ ,  $v = 100(\ln 2)^3$ .

$$100[\ln(0+k)]^3 + C = 100(\ln 2)^3$$

\*

$$100(\ln k)^3 = 100(\ln 2)^3$$

$$k = 2$$

1

\*1M for both

12. (b) (ii)  $\frac{dE}{dt}$

$$= \left[ (2^M)(\ln 2) \left( \frac{dM}{dt} \right) \right] \left[ \log \left( \frac{dv}{dt} \right) \right]^2 + (2^M) \left\{ \left[ 2 \log \left( \frac{dv}{dt} \right) \right] \left( \frac{\frac{d^2v}{dt^2}}{\frac{dv}{dt} \ln 10} \right) \right\}$$

1M + 1A

$$= (2^M)(\ln 2) \left( \frac{dM}{dt} \right) \left[ \log \left( \frac{dv}{dt} \right) \right]^2 + \left( \frac{2^{M+1}}{\ln 10} \right) \left[ \log \left( \frac{dv}{dt} \right) \right] \left( \frac{\frac{d^2v}{dt^2}}{\frac{dv}{dt}} \right)$$

$$\frac{dE}{dt} \Big|_{t=T}$$

$$= (2^{\log_2 e})(\ln 2) \left( \frac{dM}{dt} \Big|_{t=T} \right) \left[ \log \left( \frac{dv}{dt} \Big|_{t=T} \right) \right]^2 + \left( \frac{2^{\log_2 e+1}}{\ln 10} \right) \left[ \log \left( \frac{dv}{dt} \Big|_{t=T} \right) \right] \left( \frac{\frac{d^2v}{dt^2} \Big|_{t=T}}{\frac{dv}{dt} \Big|_{t=T}} \right)$$

1M

$$= (e \ln 2) \left( \frac{dM}{dt} \Big|_{t=T} \right) \left[ \log \left( \frac{300(2)^2}{e^2} \right) \right]^2$$

$$= (e \ln 2) \left( \frac{dM}{dt} \Big|_{t=T} \right) \left[ \log \left( \frac{1200}{e^2} \right) \right]^2$$

$$0 \leq \frac{dE}{dt} \Big|_{t=T} \leq \left[ \log \left( \frac{1200}{e^2} \right) \right]^2 \quad 1M$$

$$0 \leq (e \ln 2) \left( \frac{dM}{dt} \Big|_{t=T} \right) \left[ \log \left( \frac{1200}{e^2} \right) \right]^2 \leq \left[ \log \left( \frac{1200}{e^2} \right) \right]^2$$

$$0 \leq \frac{dM}{dt} \Big|_{t=T} \leq \frac{1}{e \ln 2}$$

$$0 \leq \frac{dM}{dt} \Big|_{t=T} < 1$$

$T$  s after the start of the experiment, the mass of the particle increases at a rate below 1 unit/s.

The claim is agreed.

1A f.t.

**End of Marking Scheme**